

## B Math III, Second Semester, March 09, 2007, Midterm

**Answer any six questions. All questions carry equal weight.**

1. (a). Suppose that  $X_n \xrightarrow{p} X$  and  $Y_n \xrightarrow{p} Y$ . Show that if  $f(x, y)$  is a continuous function, then

$$f(X_n, Y_n) \xrightarrow{p} f(X, Y).$$

(b). If in addition  $P(Y = 0) = 0$ , then show that

$$\frac{X_n}{Y_n} \xrightarrow{p} \frac{X}{Y}.$$

2. Suppose that  $\xi_1, \xi_2, \dots$  are independent and identically distributed random variables such that  $E[\xi_1] = \mu$  and  $E[\xi_1^2] < \infty$ . Define random variables recursively

$$X_k = \rho X_{k-1} + \xi_k, \quad k \geq 1, \quad \text{and} \quad X_0 = 0$$

where

$$|\rho| < 1.$$

Let

$$\bar{X}_n = \frac{1}{n} \sum_{k=1}^n X_k.$$

(a). Show that

$$E[\bar{X}_n] \rightarrow \frac{\mu}{1-\rho}$$

and for any  $l \geq 0$ ,

$$\text{Cov}(X_{k+l}, X_k) = \rho^l (1 + \rho^2 + \dots + \rho^{2k-2}).$$

(b). Show that

$$\bar{X}_n = \frac{1}{n} \sum_{k=1}^n X_k \xrightarrow{p} \frac{\mu}{1-\rho}.$$

3. Let  $(X_1, \dots, X_n)$  be a random sample with a common mean  $\mu$  and a common variance  $\sigma^2$ . Suppose that  $g(y)$  is differentiable such that

$$g'(\mu) \neq 0,$$

where  $g'(\mu)$  is the first derivative of  $g(y)$  at  $y = \mu$ . Recall that this means, if we let  $R(\mu, h)$  to be such that  $g(\mu + h) - g(\mu) - hg'(\mu) = hR(\mu, h)$ , then

$$\sup_{0 < |h| \leq \varepsilon} |R(\mu, h)| \rightarrow 0 \quad \text{as} \quad \varepsilon \rightarrow 0.$$

(a). Show that  $R(\mu, (\bar{X}_n - \mu)) \xrightarrow{p} 0$  and  $\sqrt{n}(\bar{X}_n - \mu) R(\mu, (\bar{X}_n - \mu)) \xrightarrow{p} 0$ .

(b). Using (a), show that  $\sqrt{n}(g(\bar{X}_n) - g(\mu))$  converges in distribution to a normal distribution with mean 0 and variance  $\sigma^2 (g'(\mu))^2$ .

(If you are applying CLT in (a) and/or (b), explain.)

(c). Let  $S_n$  be binomially distributed with parameters  $n$  and  $p$ ,  $0 < p < 1$ . Deduce from part (b) that  $\sqrt{n}\left(\left(\frac{S_n}{n}\right)^2 - p^2\right)$  converges in distribution to a normal distribution. Give the explicit expression for the variance of the limiting normal distribution.

4. (a). Show that if the random vector  $\mathbf{V} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_k - \boldsymbol{\phi}\boldsymbol{\phi}')$  where  $\mathbf{I}_k$  is a unit matrix of order  $k$  and  $\boldsymbol{\phi}' = (\sqrt{\pi_1}, \dots, \sqrt{\pi_k})$  with  $\pi_j > 0$ ,  $j = 1, \dots, k$ , and  $\pi_1 + \dots + \pi_k = 1$ , then the quadratic  $\mathbf{V}'\mathbf{V}$  has a  $\chi^2(k-1)$  distribution with  $k-1$  degrees of freedom.

(b). Consider a multinomial population with  $k$  (mutually exclusive) outcomes, with the corresponding probabilities  $\pi_1, \dots, \pi_k$ . We have a random sample of size  $n$  from this population. Let  $N_j =$  the number of times the  $j$ th outcome occurs in the sample,  $j = 1, \dots, k$ . Define

$$\chi_n^2 = \sum_{j=1}^k \frac{(N_j - n\pi_j)^2}{n\pi_j}.$$

Show that, when  $\pi_1, \dots, \pi_k$  are the true population probabilities,  $\chi_n^2$  converges in distribution to a  $\chi^2(k-1)$  distribution with  $k-1$  degrees of freedom.

5. The population is as in question 4 (b). For testing the simple null hypothesis

$$\pi_1 = \pi_{10}, \quad \pi_{k-1} = \pi_{k-1,0},$$

consider the likelihood ratio statistic

$$\lambda_n = \frac{l_n(\boldsymbol{\pi}_0)}{\sup_{\boldsymbol{\pi}} l_n(\boldsymbol{\pi})},$$

where  $\boldsymbol{\pi} = (\pi_1, \dots, \pi_k)'$  and

$$l_n(\boldsymbol{\pi}) = \pi_1^{N_1} \dots \pi_k^{N_k}$$

with  $N_j$  as defined in question 4 (b) based on a random sample of size  $n$ .

(a). Show that  $\sup_{\boldsymbol{\pi}} l_n(\boldsymbol{\pi}) = l_n(\hat{\boldsymbol{\pi}}_n)$  where  $\hat{\boldsymbol{\pi}}_n = \left(\frac{N_1}{n}, \dots, \frac{N_k}{n}\right)'$ .

(b). Show that, under the null hypothesis

$$-2 \log \lambda_n - \sum_{j=1}^k \frac{(N_j - n\pi_{j0})^2}{n\pi_{j0}} \xrightarrow{p} 0.$$

6. Consider two discrete populations with respective cell probabilities  $\boldsymbol{\pi}_1 = (\pi_{11}, \dots, \pi_{1k})$  and  $\boldsymbol{\pi}_2 = (\pi_{21}, \dots, \pi_{2k})$ . Define the column vectors

$$\boldsymbol{\phi}_1 = (\sqrt{\pi_{11}}, \dots, \sqrt{\pi_{1k}})', \quad \boldsymbol{\phi}_2 = (\sqrt{\pi_{21}}, \dots, \sqrt{\pi_{2k}})'.$$

Let the vector  $\mathbf{V}_1$  be  $k$ -variate normal with mean vector  $\mathbf{0}$  and variance-covariance matrix  $\mathbf{I}_k - \boldsymbol{\phi}_1 \boldsymbol{\phi}_1'$ . Similarly let  $\mathbf{V}_2$  be  $k$ -variate normal with mean vector  $\mathbf{0}$  and variance-covariance matrix  $\mathbf{I}_k - \boldsymbol{\phi}_2 \boldsymbol{\phi}_2'$ . Assume that  $\mathbf{V}_1$  and  $\mathbf{V}_2$  are independent.

Let

$$\mathbf{V} = \begin{bmatrix} \mathbf{V}_1 \\ \mathbf{V}_2 \end{bmatrix} = (\mathbf{V}'_1, \mathbf{V}'_2)',$$

and

$$\boldsymbol{\phi} = \begin{bmatrix} \boldsymbol{\phi}_1 & \mathbf{0} \\ \mathbf{0} & \boldsymbol{\phi}_2 \end{bmatrix}_{2k \times 2}.$$

Then

(a). Show that  $\mathbf{V}$  is  $2k$ -variate normal with mean vector  $\mathbf{0}$  and variance-covariance matrix  $\mathbf{I}_{2k} - \boldsymbol{\phi} \boldsymbol{\phi}'$ .

(b). Show that if the quadratic  $\mathbf{V}'\mathbf{C}\mathbf{V}$  is such that

$$\mathbf{C}^2 = \mathbf{C} \quad \text{and} \quad \mathbf{C}\boldsymbol{\phi} = \boldsymbol{\phi},$$

then  $\mathbf{V}'\mathbf{C}\mathbf{V}$  has a  $\chi^2(l)$  distribution with the degrees of freedom  $l$  given by

$$l = \text{Rank}(\mathbf{C}) - 2.$$

(c). Generalize the preceding two population situation to the case of  $m$  populations.

7 Consider two multinomial populations, each with the same number of outcomes or cells, with respective cell probabilities  $\boldsymbol{\pi}_1 = (\pi_{11}, \dots, \pi_{1k})$  and  $\boldsymbol{\pi}_2 = (\pi_{21}, \dots, \pi_{2k})$ . We want to test the null hypothesis

$$(\pi_{11}, \dots, \pi_{1k}) = (\pi_{21}, \dots, \pi_{2k}) = (\pi_1, \dots, \pi_k), \text{ say.}$$

Suppose that we have a sample of size  $n_1$  from the first population, and denote the corresponding sample frequencies by  $N_{11}, \dots, N_{1k}$ . Similarly, let  $N_{21}, \dots, N_{2k}$  be the sample frequencies corresponding to the sample of size  $n_2$  from the second population.

An appropriate Chi-squared statistic is

$$\chi_n^2 = \sum_{j=1}^k \frac{(N_{1j} - n_1 \hat{\pi}_j)^2}{n_1 \hat{\pi}_j} + \sum_{j=1}^k \frac{(N_{2j} - n_2 \hat{\pi}_j)^2}{n_2 \hat{\pi}_j}$$

where

$$\hat{\pi}_j = \frac{N_{1j} + N_{2j}}{n_1 + n_2}.$$

(a). Assuming for simplicity that  $n_1 = n_2 = n$ , show that  $\chi_n^2$  has asymptotically a  $\chi^2(k-1)$  with  $k-1 + k-1 - k = k-1$  degrees of freedom.

(b). Generalize the preceding parts (a) to the situation of  $m$  populations, and show that the resulting  $\chi_n^2$  statistic will have asymptotically a  $\chi^2(l)$  with  $l = (m-1)(k-1)$  degrees of freedom.